



Institut
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Less is More

Dimensionality Reduction from a Theoretical Perspective

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Notations and Model

Optimal..

..distinguisher

..dimension reduction

Comparison to..

..PCA

..LDA

Numerical Comparison

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Conclusion



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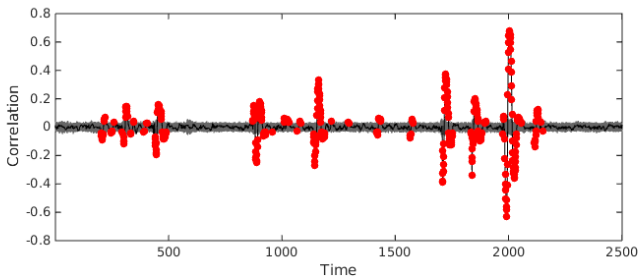
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Motivation



large number of samples/ points of interest



Motivation

Problem (*profiled* and *non-profiled* side-channel distinguisher)

How to reduce dimensionality of multi-dimensional measurements?

Motivation

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How to reduce dimensionality of multi-dimensional measurements?

Wish list

- simplification of the problem
- concentration of the information (to distinguish using fewer traces)
- improvement of the computational speed



State-of-the-Art I

Selection of points of interest

- manual selection of educated guesses [Oswald et al., 2006]
- automated techniques: sum-of-square differences (SOSD) and t-test (SOST) [Gierlichs et al., 2006]
- wavelet transforms [Debande et al., 2012]

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Leakage detection metrics

- ANOVA (e.g. [Choudary and Kuhn, 2013, Danger et al., 2014]) or [Bhasin et al., 2014] (*Normalized Inter-Class Variance (NICV)*)

State-of-the-Art II

Principal Component Analysis

- compact templates in [Arhambeau et al., 2006]
- reduce traces in [Batina et al., 2012]
- eigenvalues as a security metric [Guilley et al., 2008]
- eigenvalues as a distinguisher [Souissi et al., 2010]

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State-of-the-Art II

Linear Discriminant Analysis

- improved alternative
- takes inter-class variance and intra-class variance into account
- empirical comparisons [Standaert and Archambeau, 2008, Renauld et al., 2011, Strobel et al., 2014]

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But..

- advantages due to the statistical tools, their implementation, data set ...
- no clear rationale to prefer one method!



Contribution

- dimensional reduction in SCA from a theoretical viewpoint
- assuming attacker has full knowledge of the leakage
- derivation of the optimal dimensionality reduction

“Less is more”

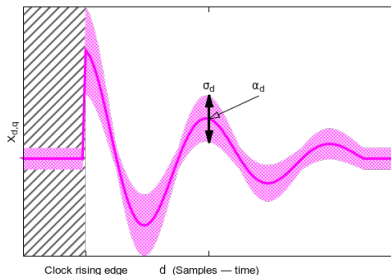
Advantages of dimensionality reduction can come with no impact on the attack success probability!

- comparison to PCA and LDA: theoretically and practically

Notations

- unknown secret key k^* , key byte hypothesis k
- D different samples, $d = 1, \dots, D$
- Q different traces/ queries, $q = 1, \dots, Q$
- matrix notation $M^{D,Q}$ (D rows, Q columns)
- leakage function φ
- sensitive variable: $Y_q(k) = \varphi(T_q \oplus k)$ (normalized variance $\forall q$)

Model



- trace $X_{d,q} = \alpha_d Y_q(k^*) + N_{d,q}$
- traces $X^{D,Q} = \alpha^D Y^Q(k^*) + N^{D,Q}$
- noise: zero-mean Gaussian distribution, covariance Σ
- independent of q but can be correlated among d



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Optimal distinguisher

Data processing theorem [Cover and Thomas, 2006]

Any preprocessing like dimensionality reduction can only decrease information.

- optimal means optimizing the success rate
- known leakage model: optimal attack \Rightarrow template attack
- maximum likelihood principle

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- Given:
 - Q traces of dimensionality D in a matrix $x^{D,Q}$
 - for each trace x_q^D : a plaintext/ciphertext t_q

Optimal distinguisher

$$\begin{aligned}\mathcal{D}(x^{D,Q}, t^Q) &= \arg \max_k p(x^{D,Q} | t^Q, k^* = k) \\ &= \arg \max_k p_{N^{D,Q}}(x^{D,Q} - \alpha^D y^Q(k)) \\ &= \arg \max_k \prod_{q=1}^Q p_{N_q^D}(x_q^D - \alpha^D y_q(k))\end{aligned}$$

where

$$p_{N_q^D}(z^D) = \frac{1}{\sqrt{(2\pi)^D |\det \Sigma|}} \exp\left(-\frac{1}{2}(z^D)^\top \Sigma^{-1} z^D\right).$$

Optimal dimension reduction

Theorem

The optimal attack on the multivariate traces $x^{D,Q}$ is equivalent to the optimal attack on the monivariate traces \tilde{x}^Q , obtained from $x^{D,Q}$ by the formula:

$$\tilde{x}_q = (\alpha^D)^\top \Sigma^{-1} x_q^D \quad (q = 1, \dots, Q).$$

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$$\tilde{x}_q = (\alpha^D)^T \Sigma^{-1} x_q^D \quad (q = 1, \dots, Q).$$

scalar = column $D \cdot D \times D \cdot$ row D

Proof I

- taking the logarithm, the optimal distinguisher $\mathcal{D}(x^{D,Q}, t^Q)$ rewrites

$$\mathcal{D}(x^{D,Q}, t^Q) = \arg \min_k \sum_{q=1}^Q (x_q^D - \alpha^D y_q(k))^T \Sigma^{-1} (x_q^D - \alpha^D y_q(k)) .$$

Proof I

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- expansion gives

$$\underbrace{(x_q^D)^T \Sigma^{-1} x_q^D}_{\text{cst. } C \text{ independent of } k} - 2(\alpha^D)^T y_q(k) \Sigma^{-1} x_q^D + (y_q(k))^2 (\alpha^D)^T \Sigma^{-1} \alpha^D$$

cst. C independent of k

$$= C - 2y_q(k) [(\alpha^D)^T \Sigma^{-1} x_q^D] + (y_q(k))^2 [(\alpha^D)^T \Sigma^{-1} \alpha^D]$$

$$= [(\alpha^D)^T \Sigma^{-1} \alpha^D] \left(y_q(k) - \frac{(\alpha^D)^T \Sigma^{-1} x_q^D}{(\alpha^D)^T \Sigma^{-1} \alpha^D} \right)^2 + C'.$$

Proof II

- so, for $\mathcal{D}(x^{D,Q}, t^Q)$ we obtain

$$\begin{aligned}\mathcal{D}(x^{D,Q}, t^Q) &= \arg \min_k \sum_{q=1}^Q \left(y_q(k) - \frac{(\alpha^D)^\top \Sigma^{-1} x_q^D}{(\alpha^D)^\top \Sigma^{-1} \alpha^D} \right)^2 [(\alpha^D)^\top \Sigma^{-1} \alpha^D] \\ &= \arg \min_k \sum_{q=1}^Q \frac{(\tilde{x}_q - y_q(k))^2}{\tilde{\sigma}^2},\end{aligned}$$

where

$$\begin{cases} \tilde{x}_q &= \tilde{\sigma}^2 \cdot (\alpha^D)^\top \Sigma^{-1} x_q^D, \\ \tilde{\sigma} &= ((\alpha^D)^\top \Sigma^{-1} \alpha^D)^{-1/2}. \end{cases}$$

Discussion

Optimal dimension reduction

Optimal distinguisher can be computed either:

- on multivariate traces x_q^D , with a noise covariance matrix Σ
- on monivariate traces \tilde{x}_q , with scalar noise of variance $\tilde{\sigma}^2$.

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-
- optimal dimensionality reduction does not depend on the distribution of $Y^D(k)$
 - also not on the *confusion coefficient* [Fei et al., 2012]
 - only on the signal weights α^D and on the noise covariance Σ

Corollary

After optimal dimensionality reduction, the signal-noise-ratio is given by

$$\frac{1}{\tilde{\sigma}^2} = (\alpha^D)^\top \Sigma^{-1} \alpha^D.$$

In the paper...

Examples

- white noise:

$$\widetilde{\text{SNR}} = \sum_{d=1}^D \text{SNR}_d$$

- autoregressive noise (confirmed on dpacontest v2)

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Dimensionality Reduction from a Theoretical Perspective

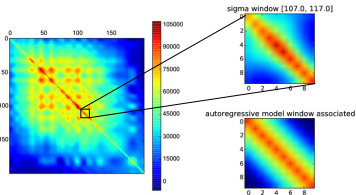
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Thibaut Marlot⁴, and Olivier Riou^{1,2}

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² UPMC Sorbonne Universités, UFR Sorbonne Sciences, Paris, France
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Abstract. Reduced the dimensionality of the observations in the source problem, an auto-correlated matrix, is shown to improve the performance of the source problem. The auto-correlation of the source observations is shown to be related to the performance of the source problem. The auto-correlation of the source observations is shown to be related to the performance of the source problem. The auto-correlation of the source observations is shown to be related to the performance of the source problem.

1 Introduction

High-dimensional signals are often observed in many applications. Typical examples are signals in communications systems or images. Typical examples are signals in communications systems or images. Typical examples are signals in communications systems or images. Typical examples are signals in communications systems or images. Typical examples are signals in communications systems or images.





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Comparison to PCA

Classical PCA

- centered data $M_{d,q} = X_{d,q} - \frac{1}{Q} \sum_{q'=1}^Q X_{d,q'}$ ($1 \leq q \leq Q, 1 \leq d \leq D$)
- directions of PCA: eigenvectors of $M^{D,Q}(M^{D,Q})^T$
- drawback: depends both on data and noise

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Inter-class PCA [Archambeau et al., 2006]

- centered column $\frac{1}{\sum_{1 \leq q \leq Q} 1_{Y_q=y}} \sum_{1 \leq q \leq Q} 1_{Y_q=y} X_q^D$
- takes into account the sensitive variable Y
- noise is averaged away

Comparison to PCA

For classical PCA

Asymptotically as $Q \rightarrow +\infty$,

$$\frac{1}{Q} M^{D,Q} (M^{D,Q})^T \rightarrow \alpha^D (\alpha^D)^T + \Sigma.$$



Eigenvectors?

Comparison to PCA

For classical PCA

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Eigenvectors?

Proposition

Asymptotically, Inter-class PCA has only one principal direction, namely the vector α^D .

Comparison to PCA

Proposition

The asymptotic SNR after projection using Inter-class PCA is equal to

$$\frac{\|\alpha^D\|_2^4}{(\alpha^D)^T \Sigma \alpha^D}.$$

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Theorem

The SNR of the asymptotic Inter-class PCA is smaller than the SNR of the optimal dimensionality reduction.

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Corollary

The asymptotic Inter-class PCA has the same SNR as the optimal dimensionality reduction if and only if α^D is an eigenvector of Σ . In this case, both dimensionality reductions are equivalent.

Comparison to LDA

- computes the eigenvectors of $S_w^{-1}S_b$
- S_w is the *intra-class scatter matrix*, asymptotically equal to Σ
- S_b is the *inter-class scatter matrix*, equal to $\alpha^D(\alpha^D)^T$.

Proposition

Asymptotically, LDA has only one principal direction, namely the vector $\Sigma^{-1}\alpha^D$.

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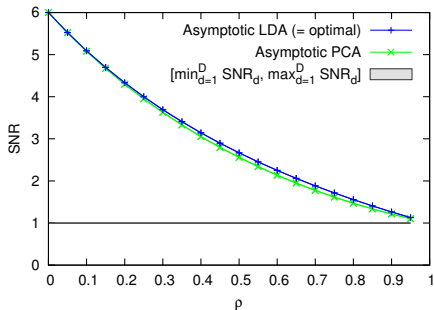
The asymptotic LDA computes exactly the optimal dimensionality reduction.

Asymptotic PCA and LDA

- $D = 6$ for autoregressive noise with $\sigma = 1$ and different ρ

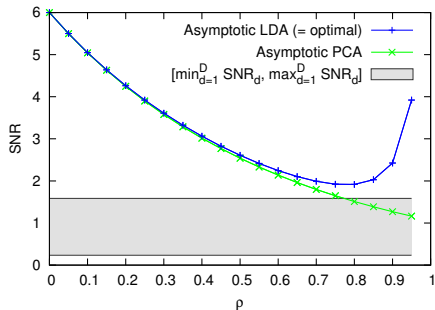
(a) Equal $\text{SNR}_d = 1, 1 \leq d \leq D$

$$\alpha^D = (1, 1, 1, 1, 1, 1)^T$$



(b) Varying $\text{SNR}_d, 1 \leq d \leq D$

$$\alpha^D = \sqrt{6.0/6.4} \cdot (1.0, 1.1, 1.2, 1.3, 0.9, 0.5)^T$$





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Practical Validation

- DPA CONTEST V2, one clock cycle $D = 200$
- normalized Hamming weight
- precharacterization of the model parameter α^D and Σ (details in the paper)

- $\max_{d=1}^D \hat{\alpha}_d^2 / \hat{\Sigma}_{d,d} = 1.69 \cdot 10^{-3}$ (no dimensionality reduction)
- $\text{SNR}_{\text{PCA}} = \frac{((\hat{\alpha}^D)^\top \hat{\alpha}^D)^2}{(\hat{\alpha}^D)^\top \hat{\Sigma} \hat{\alpha}^D} = 1.36 \cdot 10^{-3}$ (PCA)
- $\text{SNR}_{\text{LDA}} = (\hat{\alpha}^D)^\top \hat{\Sigma} \hat{\alpha}^D = 12.78 \cdot 10^{-3}$ (LDA)



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Conclusion and Perspectives

Optimal dimension reduction...

- is part of the optimal attack
- can be achieved *without* losing success probability



Conclusion and Perspectives

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- LDA asymptotically achieves the same projection as optimal
- when weakly correlated (Σ is identity matrix) PCA is nearly equivalent to optimal/ LDA



Conclusion and Perspectives

Optimal dimension reduction...

- is part of the optimal attack
- can be achieved *without* losing success probability

- LDA asymptotically achieves the same projection as optimal
- when weakly correlated (Σ is identity matrix) PCA is nearly equivalent to optimal/ LDA
- ★ extend to non-Gaussian noise
- ★ comparison to machine-learning techniques





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